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## Global Warming Effects on U.S. Hurricane Damage

Kerry Emanuel

Program in Atmospheres, Oceans, and Climate

Room 54-1814, Massachusetts Institute of Technology

77 Massachusetts Avenue

Cambridge, Massachusetts, 02139

(617) 253-2462

[emanuel@mit.edu](mailto:emanuel@mit.edu)

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**Abstract**

While many studies of the effects of global warming on hurricanes predict an increase in various metrics of Atlantic basin-wide activity, it is less clear that this signal will emerge from background noise in measures of hurricane damage, which depend largely on rare, high intensity landfalling events and are thus highly volatile compared to basin-wide storm metrics. Using a recently developed hurricane synthesizer driven by large-scale meteorological variables derived from global climate models, we generate 1000 artificial 100-year time series of Atlantic hurricanes that make landfall along the U.S. Gulf and east coasts, for four climate models and for current climate conditions as well as for the warmer climate of 100 years hence under IPCC emissions scenario A1b. These synthetic hurricanes damage a portfolio of insured property, according to an aggregate wind-damage function; damage from flooding is not considered here. Assuming that the hurricane climate changes linearly with time, we create a 1000-member ensemble of time series of property damage. Three of the four climate models used produce increasing damage with time, with the global warming signal emerging on time scales of 40, 113, and 170 years, respectively. We point out, however, that probabilities of damage increase significantly well before such emergence time scales and show that probability density distributions of aggregate damage become appreciably separated from those of the control climate on time scales as short as 25 years. For the fourth climate model, damages decrease with time, but the signal is weak.

34

35 **1. Introduction**

36 Several lines of evidence suggest that anthropogenic climate change may have a substantial  
37 influence on tropical cyclone activity around the world. Global warming generally increases the  
38 thermodynamic potential for tropical cyclones (Emanuel, 1987) while changing atmospheric circulation,  
39 humidity, and other factors affect both the probability of genesis and the subsequent evolution of the  
40 storms (Emanuel, 2007, Vecchi and Soden, 2007). Strictly thermodynamic considerations lead to the  
41 expectation that, globally, tropical cyclone frequency should diminish, but the incidence of high intensity  
42 events should increase (Emanuel et al., 2008). There is already evidence that the fraction of high  
43 intensity storms is indeed increasing (Webster et al., 2005, Elsner et al., 2008), although the global total  
44 number of tropical cyclones has not so far exhibited any significant trend (Emanuel, 2005). In the North  
45 Atlantic region, where tropical cyclone records are longer and generally of better quality than  
46 elsewhere, power dissipation by tropical cyclones is highly correlated with sea surface temperature  
47 during hurricane season in the regions where storms typically develop (Emanuel, 2005) and with the  
48 difference between the local sea surface temperature and the tropical mean sea surface temperature  
49 (Swanson, 2008, Vecchi et al., 2008). The author has argued that in the North Atlantic region, the  
50 decadal variations of the sea surface temperature itself appear to be driven mostly by anthropogenic  
51 changes in greenhouse gases and aerosols (Mann and Emanuel, 2006, Emanuel, 2008). Other studies  
52 hold attribution of past changes in tropical cyclone activity to anthropogenic climate change to be  
53 equivocal (Knutson et al., 2010).

54 Quite a few attempts have been made to use global climate models to make projections of the  
55 response of tropical cyclones to global warming. One method simply detects explicitly simulated storms  
56 in the models and notes how their levels of activity change with climate. This approach has been taken

57 by numerous groups (e.g. Bengtsson et al., 1996, Sugi et al., 2002, Oouchi et al., 2006, Yoshimura et al.,  
58 2006, Bengtsson et al., 2007) and is becoming more popular as the horizontal resolution of global  
59 climate models improves. But even horizontal grid spacing as low as 20 km (Oouchi et al., 2006) cannot  
60 resolve the critical eyewall region of the cyclones, and invariably the maximum wind speed of simulated  
61 storms is truncated at relatively low values by the lack of horizontal resolution (Zhao et al., 2009).  
62 Recent work by Rotunno et al. (2009) suggests that horizontal grid spacing of less than 1 km is needed to  
63 properly resolve intense storms. A second approach to the problem is to use statistical relationships  
64 between tropical cyclones and large-scale predictors to estimate tropical cyclone activity as a function of  
65 variables that are resolved by climate models (e.g. Camargo et al., 2007a, Camargo et al., 2007b). One  
66 potential drawback of this approach is that the statistics are trained largely on natural variability, much  
67 of which is regional; it is not clear that such indices will perform well when applied to global climate  
68 change. Another approach to quantifying the relationship between climate and tropical cyclone activity  
69 is to “downscale” tropical cyclone activity from reanalysis or climate model data sets, as pioneered by  
70 Knutson et al. (2007) and Emanuel et al. (2008). Such techniques involve running high-resolution,  
71 detailed models capable of resolving tropical cyclones, using boundary conditions supplied by reanalysis  
72 or climate model data sets. This combines the advantage of relatively robust estimates of large-scale  
73 conditions by the reanalyses or climate models with the high fidelity simulation of tropical cyclones by  
74 the embedded high-resolution models. As shown by Knutson et al. (2007) and Emanuel et al. (2008),  
75 these techniques are remarkably successful in reproducing observed tropical cyclone climatology in the  
76 period 1980-2006, particularly in the North Atlantic region, when driven by NCAR-NCEP reanalysis data  
77 (Kalnay and co-authors, 1996). A recent comparison between storms produced explicitly in a high-  
78 resolution global simulation and those downscaled from the same global model (Emanuel et al., 2010)  
79 confirms that even at grid spacing of 14 km, global models truncate the important high-intensity end of

80 the spectrum of tropical cyclones, and reveals substantial differences between the explicit and  
81 downscaled storm activity.

82 The general conclusion from all of these studies is that while the global frequency of tropical  
83 cyclones is likely to diminish, the frequency of high-intensity events will probably increase as the planet  
84 continues to warm (Knutson et al., 2010). Since most wind-related damage is owing to high intensity  
85 events, this would imply an increase in wind damage. On the other hand, there is large regional and  
86 model-to-model variability in projections of climate change effects on tropical cyclones, so confidence in  
87 any regional projections must be correspondingly low. For the North Atlantic, a downscaling of a set of  
88 global climate model projections shows that five out of seven of the models predict substantial  
89 increases in power dissipation over the 21<sup>st</sup> century (Emanuel et al., 2008), while a recent downscaling  
90 using a comprehensive tropical cyclone model run on a 9-km mesh shows a near doubling of the  
91 frequency of high intensity events (Bender et al., 2010). Thus the weight of current evidence suggests a  
92 possibly substantial increase in damaging Atlantic hurricanes over the current century, though  
93 uncertainty remains large.

94 While basin-wide metrics of tropical cyclone activity show statistically robust changes in the  
95 aforementioned model-based projections and may already be evident in observations, there is little  
96 evidence for a trend in tropical cyclone-related damage in the U.S (Pielke et al., 2008). This is not  
97 surprising, as most wind-related damage is done by tropical cyclones that happen to be at high intensity  
98 at the time they make landfall. This is a small subset of all storms over a relatively small fraction of their  
99 typical life spans, thus the statistical base of potentially damaging events is small compared to that of  
100 the basin-wide set of storms. These findings beg the question of how long it would take for any climate  
101 change signal to emerge from background natural variability in damage statistics. This question was  
102 addressed recently by Crompton et al. (2011), who concluded that it will take between 120 and 550

103 years for such a signal to emerge in U.S. tropical cyclone losses. They arrive at their findings using a  
104 relationship between normalized losses and Saffir-Simpson category derived from past events, and  
105 applying that relationship to projected changes in Atlantic storm activity assuming that fractional  
106 changes in the frequency of landfalling events in each category are the same as that of all storms in the  
107 North Atlantic. Thus the technique cannot account for possible changes in the tracks of storms, which  
108 may change the fraction of all events that make landfall as well as the specific locations of landfall.  
109 Moreover, quantization of storms into only five categories, with most of the damage being done by  
110 storms of the highest three categories, may alter the signal because it misses changes within categories.  
111 (For example, an increase of intensity within category 5, which is open-ended, would presumably cause  
112 increased damage even if the number of landfalling category 5 storms remained constant.) These  
113 limitations motivate the present study, which applies somewhat different methods to the problem, as  
114 described in the next section.

## 115 **2. Method**

116 Here we apply the tropical cyclone downscaling technique of Emanuel et al. (2008). Briefly, this  
117 method embeds a specialized, atmosphere-ocean coupled tropical intensity model in the large-scale  
118 atmosphere-ocean environment represented by the global climate model data. The tropical cyclone  
119 model is initialized from weak, warm-core vortices seeded randomly in space and time, and whose  
120 movement is determined with a beta-and-advection model driven by the flows derived from the climate  
121 model daily wind fields. The thermodynamic state used by the intensity model is derived from monthly  
122 mean climate model data together with current climatological estimates of ocean mixed layer depths  
123 and sub-mixed layer thermal stratification; these ocean parameters are held fixed at their current  
124 climatological values. Wind shear, used as input to the intensity model, is likewise derived from the  
125 climate model wind fields. In practice, a large proportion of the initial seeds fail to amplify to at least

126 tropical storm strength and are discarded; the survivors are regarded as constituting an estimate of the  
127 tropical cyclone climatology for the given climate state.

128           We apply this method to each of four climate models, applying enough seeds to produce 5000  
129 U.S. landfalling storms in each case. The models are the CNRM-CM3 model of the Centre National de  
130 Recherches Météorologiques, Météo-France; the ECHAM5 model of the Max Planck Institution; the  
131 GFDL-CM2.0 model of the NOAA Geophysical Fluids Dynamics Laboratory; and the MIROC 3.2 model of  
132 the CCSR/NIES/FRCGC, Japan. We chose these four models from the set of seven models used by  
133 Emanuel et al. (2008) which were in turn selected based on the availability of model output needed by  
134 the downscaling technique; the four used here are broadly representative of the larger set of seven. We  
135 apply the downscaling to simulations of the 20<sup>th</sup> century climate, using model output for the period  
136 1981-2000, and to simulations of a warming climate under IPCC scenario SRES A1b, using model output  
137 from the period 2081-2100. We use these synthetic events to construct 1000 stationary time series of  
138 100 years length each, representing the climate averaged over 1981-2000 and over 2081-2100,  
139 respectively. We construct the time series by randomly drawing, each year, from a Poisson distribution  
140 based on the overall annual frequency of events in the set. The 1000-member ensemble is created by  
141 repeating the process using different random draws from the Poisson distribution each year. Thus we  
142 have 1000 100-year time stationary series for each of four climate models for each of two time frames.  
143 As detailed below, we will use these time series to create 1000 100-year time series of damage in an  
144 evolving climate by linearly combining the two stationary time series sets with a weighting that varies  
145 linearly in time.

146           We note here that the intent is to distinguish climate change-induced trends from background  
147 short-term random variability. We do not attempt to account for other sources of systematic climate  
148 change, such as changing solar or volcanic activity, or long-period natural variability, such as the Atlantic  
149 Multidecadal Oscillation (AMO); in this respect we use the same assumptions as did Crompton et al.



150 (2011). We simply note that detecting any signal of anthropogenic climate change, not just one that  
151 might be present in tropical cyclone statistics, requires one to account for other forced changes, and  
152 that long-period natural variability, such as the AMO, cannot meaningfully be considered noise in this  
153 context. If it exists and is important, its influence can, in principle, be quantified and accounted for; if it  
154 cannot be quantified then one must give up on any exercise in climate change attribution on these time  
155 scales.

156  
157         Next, we allow the simulated storms to interact with a portfolio of insured property: the  
158 Industry Exposure Database, produced by Risk Management Solutions Inc<sup>1</sup>. This consists of estimates of  
159 total insured values for each zip code and county in the U.S. and for each postcode in Europe, using  
160 sampled company premium information, census demographics and economics data, building square  
161 footage data, and representative policy terms and conditions. These total insured values and other  
162 variables are then aggregated into 100 zones distributed along the U.S. Gulf and east coasts, whose  
163 locations are shown in Figure 1<sup>2</sup>. These locations represent roughly the population-weighted  
164 geographical centers of the zones. For simplicity, we model the damage in a given zone according to the  
165 wind experienced at the position of the zone center. For each tropical cyclone event, we use a wind  
166 damage function, described presently, to estimate the fractional loss of value at in each zone, and  
167 multiply this by the total insured value of property in that zone. This gives an estimate of the total  
168 amount of damage in U.S dollars caused by each event in each zone; the total insured damage from an  
169 event is then the sum of this quantity over all zones. A drawback of this approach is that aggregating the  
170 building values into zones represented by points will make the damage more volatile than it should be,  
171 as some strong storms will pass between the zone centers and do little damage there, whereas in reality

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<sup>1</sup> [http://www.rms.com/Catastrophe/Models/us\\_industry\\_exposure.asp](http://www.rms.com/Catastrophe/Models/us_industry_exposure.asp)

<sup>2</sup> The original data is proprietary, and this aggregation was done by the provider to allow it to be used for the present purpose.

172 some of the insured property between zone centers will experience high winds. This increased volatility  
173 will decrease the climate signal-to-noise ratio and make the damage probability density functions  
174 broader than they ought to be.

175 Property damage from wind storms is observed to increase quite rapidly with wind speed.  
176 Empirical studies relating wind to damage suggest a high power-law dependence of damage on wind  
177 speed (Pielke, 2007) . For example, Nordhaus (2010) estimates that damage varies as the ninth power of  
178 wind speed for wind damage in the U.S. In reality, most structures in the U.S. and many other countries  
179 are built to withstand frequently encountered winds; it is highly unlikely, for example, that a wind of 20  
180 knots would do any damage at all. Thus we consider a damage function that produces positive values  
181 only for winds speeds in excess of a specified threshold. On physical grounds, we expect that damage  
182 should vary as the cube of the wind speed over a threshold value. Finally, we require that the fraction of  
183 the property damaged approach unity at very high wind speeds; in any event, we cannot allow it to  
184 exceed unity. A plausible function that meets these requirements is

$$185 \quad f = \frac{v_n^3}{1 + v_n^3}, \quad (1)$$

186 where  $f$  is the fraction of the property value lost and

$$187 \quad v_n \equiv \frac{\text{MAX}[(V - V_{thresh}), 0]}{V_{half} - V_{thresh}},$$

188 with  $V$  the wind speed,  $V_{thresh}$  is the wind speed at and below which no damage occurs, and  $V_{half}$  is the  
189 wind speed at which half the property value is lost. This function is plotted in Figure 2, for  
190  $V_{thresh} = 50 \text{ kts}$  and two values of  $V_{half}$  . These functions are highly idealized; in reality, property damage  
191 depends on much more than the peak wind speed experienced during a storm; for example, the

192 direction of the wind, its degree of gustiness, and the duration of damaging winds all influence the  
193 amount of damage. Also, we do not consider damage from freshwater flooding or storm surge, though  
194 the latter is, to some extent, also a nonlinear function of wind speed. The damage functions illustrated  
195 in Figure 2 can be compared to damage functions derived from theory and from insurance claims data  
196 as reviewed by Watson and Johnson (2004). By varying  $V_{half}$ , we will explore the sensitivity of the  
197 results to the damage function.

198         Using this damage function and the property values from the Industry Exposure Database in  
199 conjunction with the two synthetic tropical cyclone events sets for each of the four models, we derive  
200 1000 100-year time series of U.S insured property damage for each model and for each of the two  
201 climates considered. Then, for each ensemble member, we blend the two 100-year time series  
202 representing the two climates into a single time series by linearly combining the two damage amounts  
203 assuming that the transition from one climate to the other occurs linearly over 100 years; thus the  
204 damage in each year is a weighted average of the damage from each climate state, with the weight  
205 varying linearly with time. This results in 1000 100-year time series for each of the 4 models, each  
206 representing a transitioning climate. We use these time series to evaluate the emergence of global  
207 warming signals. As in the work of Crompton et al. (2011), we assume that the noise against which the  
208 global warming signal is measured is random variability on time scales ranging up to a few years  
209 (including, for example, ENSO-related variability) but not natural multidecadal variability. We also  
210 consider that changes in damage owing to changing distributions of property and property value are  
211 quantifiable after-the-fact and thus do not constitute noise in the system.

### 212         **3. Results**

213         Figure 3 shows property damage each year for a single ensemble member of the GFDL 20<sup>th</sup> century  
214 climate simulation, using the damage function (1) with  $V_{half} = 150 \text{ kts}$ . As expected, damage is highly

215 volatile, ranging from a few million to 275 million dollars. Figure 4 again shows a single randomly chosen  
216 ensemble member of property damage but for the blended time series in which the climate state  
217 transitions linearly from its late 20<sup>th</sup> century condition to its late 21<sup>st</sup> century condition under IPCC  
218 emissions scenario A1b. In this particular case, an upward trend in damage seems evident, though the  
219 trend is not large compared to the interannual variance. Figure 5 presents a different metric: damage  
220 accumulated over every year from 2000 to the year on the x axis, for a single ensemble member in each  
221 of the constant 20<sup>th</sup> century climate and the transitioning climate. (The standard deviations up and  
222 down from the ensemble means are also shown for comparison.) Accumulating the damage has the  
223 effect of smoothing over interannual variations and the difference in the trends becomes clear after a  
224 few decades.

225 The probability densities of damage, derived using the 1000-member ensemble of accumulated  
226 damage at various times, are shown for all four models in Figure 6. In most cases, the probability  
227 densities of the warming climate are distinct from those of the current climate by 100 years out. Note  
228 that in the case of the MIROC model, the probability densities shift toward lower damage amounts as  
229 the climate warms.

230 We also calculate the time scale over which the global warming signal may be considered to have  
231 emerged in time series of damage. Following Bender et al. (2010) and Crompton et al. (2011), we define  
232 the emergence time scale as that time after which fewer than 5% of the linear regression slopes of  
233 damage up to that time, amongst the 1000-member ensemble, are negative. In the case of the GFDL  
234 CM2.0-derived damage time series, this occurs at 40 years. In the other cases, it does not occur within  
235 the 100-year time frame of the time series. In these cases, we artificially extend the time series of each  
236 of the 1000-members of the ensemble by extrapolating the linear damage trend forward another 100  
237 years. Figure 7 shows an example of the fraction of regression slopes that are negative, as a function of

238 the length of time over which the regression is carried out. It can be shown that if the interannual  
239 variance of property damage (the “noise”) is Gaussian, and the underlying trend is linear, then the  
240 fraction of negative slopes as a function of the length of the series should be a cumulative distribution  
241 function (cdf) formed from a normal distribution; for this reason we also show a fit of such a cdf to the  
242 data. With the time series extrapolated out to 200 years, the global warming signal in property damage  
243 emerges from background noise in 113 years for the CNRM model and 170 years for the ECHAM 5  
244 model. The negative signal present in the MIROC model does not emerge within the 200-year time  
245 frame.

246 How sensitive are these results to the damage function used? As a first step in addressing this issue,  
247 we repeated the analysis using  $V_{half} = 110 \text{ kts}$  in (1), as illustrated by the dashed curve in Figure 2.  
248 Figure 8 shows the result for the CNRM simulation; this should be compared to Figure 6c. Aside from the  
249 obvious increase in the magnitude of the damage, the shape of the probability distributions and their  
250 separation with climate change is hardly distinguishable. The emergence time scale remains identical at  
251 113 years. This lack of sensitivity to details of the damage function pertains to the other models as well.

#### 252 **4. Discussion**

253 For the three global climate models that produce increasing damage in the U.S., the time scales for  
254 trends in damage to emerge from background noise range from 40 to 170 years, somewhat shorter than  
255 those reported in Crompton et al. (2011). There are several potential reasons for this, including our use  
256 of a different suite of global models and that fact that Crompton et al. (2011) assumed a proportionality  
257 between basin-wide and landfalling activity and estimated changes in damages by changes in the  
258 distribution of events within the limited 5-bin Saffir-Simpson categorization.

259 We caution that the question of when a statistically robust trend can be detected in damage time  
260 series should not be confused with the question of when climate-induced changes in damage become a  
261 significant consideration. Policies and other actions that address U.S. hurricane damage on the time  
262 scale of decades would surely distinguish the probabilistic outcome represented by, say, the 25 year  
263 probability density of a warming climate given in Figure 6a from that of the steady climate at the same  
264 lead time. Thus if climate change effects are anticipated, or detected in basin-wide storm statistics,  
265 sensible policy decisions should depend on the projected overall shift in the probability of damage  
266 rather than on a high-threshold criterion for trend emergence. This is particularly important in view of  
267 evidence that suggests that an anthropogenic climate change signal has already emerged in Atlantic  
268 hurricane records (Mann and Emanuel, 2006).

269 A number of caveats apply to the present analysis. First, we have held constant the distribution and  
270 value of insured property, not accounting for changing demographics or adaptation strategies that  
271 might reduce vulnerability to damage. We do not consider the effects of rising sea-level, which would  
272 increase vulnerability to damage by storm surges. Nor have we taken into account any changing  
273 incidence of freshwater flooding stemming from tropical cyclone rainfall. We have relied for this study  
274 on a single projected emissions scenario, SRES A1b, and the results obviously depend rather sensitively  
275 on the global climate model used to drive the downscaling. At the same time, the downscaling method  
276 itself is, no doubt, an imperfect measure of the tropical cyclone climatology that would attend a  
277 particular climate state.

## 278 **5. Summary**

279 We used a synthetic tropical cyclone generator to produce 1000 artificial time series of U.S.  
280 landfalling Atlantic hurricanes, each of 100 years length for the climate of the late 20<sup>th</sup> century and for  
281 the late 21<sup>st</sup> century, using four climate models. Some of the tropical cyclones affect properties

282 contained in a portfolio of insured property, and a damage function was used to predict how much  
283 damage each storm would do to these properties. This results in two 1000-member ensembles of 100-  
284 year times series of property damage for each of the four models: one for the climate of the late 20<sup>th</sup>  
285 century and one for the climate of the late 21<sup>st</sup> century under IPCC emissions scenario SRES A1b. These  
286 two were blended together, assuming a linear variation of climate over the 21<sup>st</sup> century, to create time  
287 series of property damage representing a transitioning climate. From these times series one can make  
288 inferences concerning the effect of anthropogenic climate change on U.S. hurricane wind-related  
289 property damage that also account for the high level of background noise inherent in the volatile  
290 statistics of intense landfalling tropical cyclones.

291 For three of the four climate models downscaled, damages increase as a result of projected global  
292 warming, but the fourth model shows a small decrease of damage with time. For the three climate  
293 models that have increasing damage, the climate change signal emerges from background variability,  
294 according to a recently published criterion, on time scales of 40, 113, and 170 years, respectively; the  
295 decreasing signal of the fourth model is not clearly distinguishable from noise even after 200 years. On  
296 the other hand, the probability distributions of damage in a warming climate become distinguished  
297 from those of background climate in as little as 25 years, thus we argue that those concerned with  
298 future U.S. country-wide tropical cyclone damage on decadal time scales would be well advised to  
299 include climate change as a consideration.

300

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### Figure Captions

373 **Figure 1:** Locations of zone centers (blue dots) used for estimating hurricane damage.

374 **Figure 2:** Fraction of property value lost as a function of winds speed using equation (1) with

375  $V_{thresh} = 50 \text{ kts}$  and  $V_{half} = 150 \text{ kts}$  (solid) and  $V_{half} = 110 \text{ kts}$  (dashed).

376 **Figure 3:** Property damage (\$ millions U.S.) each year for a single ensemble member of the GFDL CM2.0

377 model with climate held fixed at its 1981-2000 mean condition.

378 **Figure 4:** Same a Figure 3 but for a climate transitioning linearly from its state at the end of the 20<sup>th</sup>

379 century to its state at the end of the 21<sup>st</sup> century.

380 Figure 5: Accumulated damage from 2000 to the year on the abscissa, from the same two ensemble

381 members presented respectively in Figures 3 and 4. The error bars shows one standard deviation up

382 and down from the ensemble mean.

383 **Figure 6:** Probability density of accumulated property damage, across the 1000-member ensemble at

384 various times as indicated, for the GFDL CM2.0 model (a), The ECHAM5 model (b), the CNRM model (c)

385 and the MIROC model (d). Blue curves indicate constant 20<sup>th</sup> century climate, while the red curves

386 shows results for the warming climate.

387 **Figure 7:** Percentage of ensembles members with negative linear regression slopes, as a function of the

388 ending time of the time series, for the CNRM damage projections (blue). The red curve represents a fit

389 to the data of a cumulative distribution function based on a normal distribution. The emergence time

390 scale is defined as the time after which negative slopes constitute less than 5% of the total; in this case,

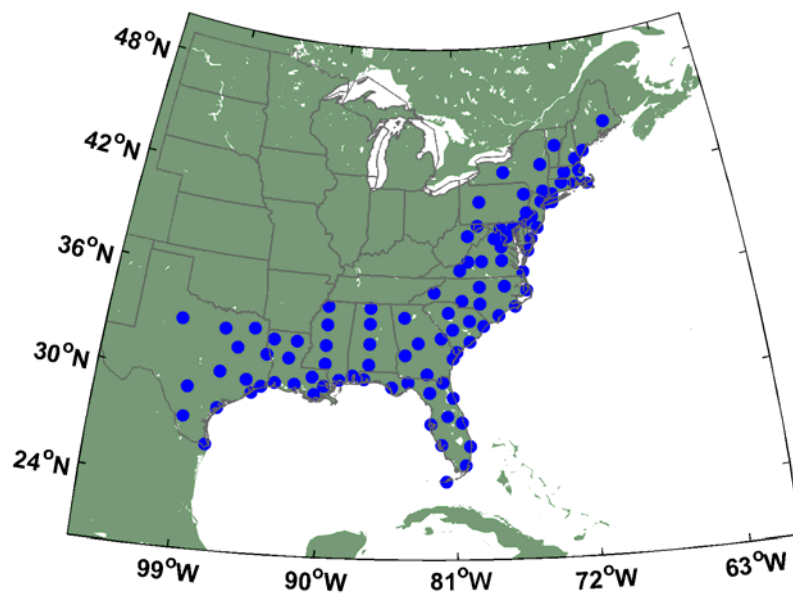
391 this occurs in year 2113.

392 **Figure 8:** Same as Figure 6c but using the damage function given by (1) using  $V_{half} = 110 \text{ kts}$ .

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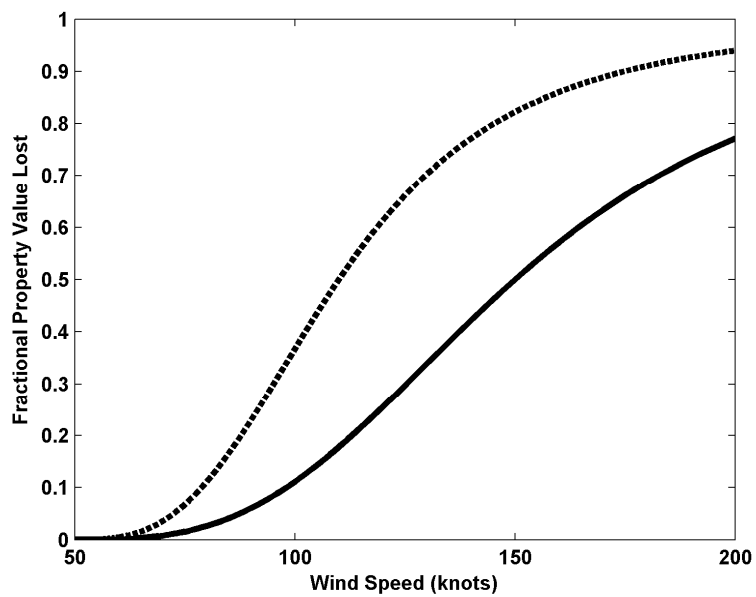
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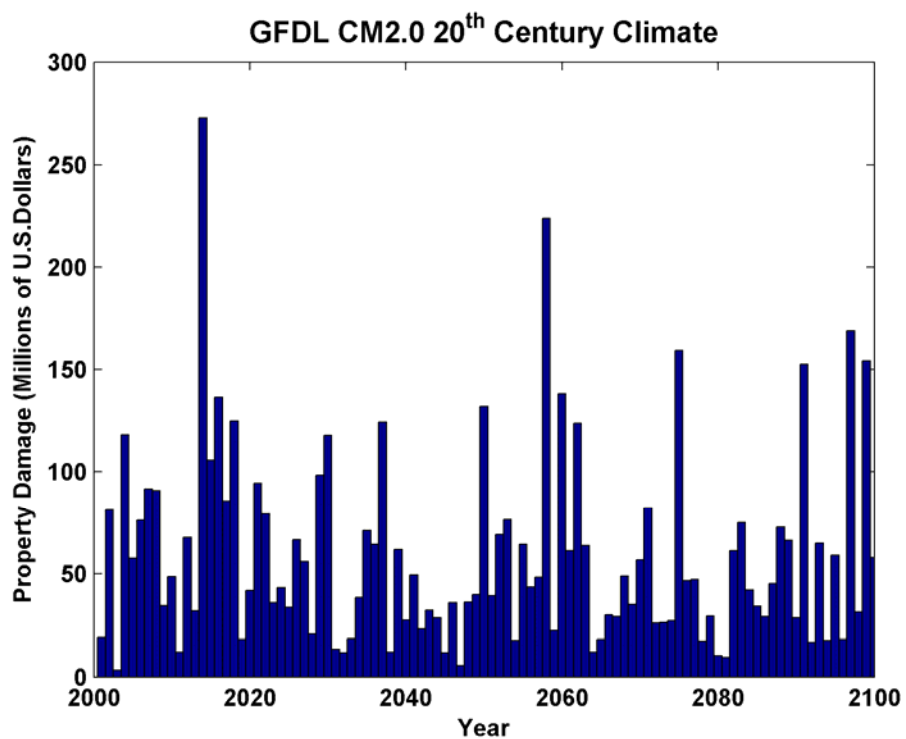
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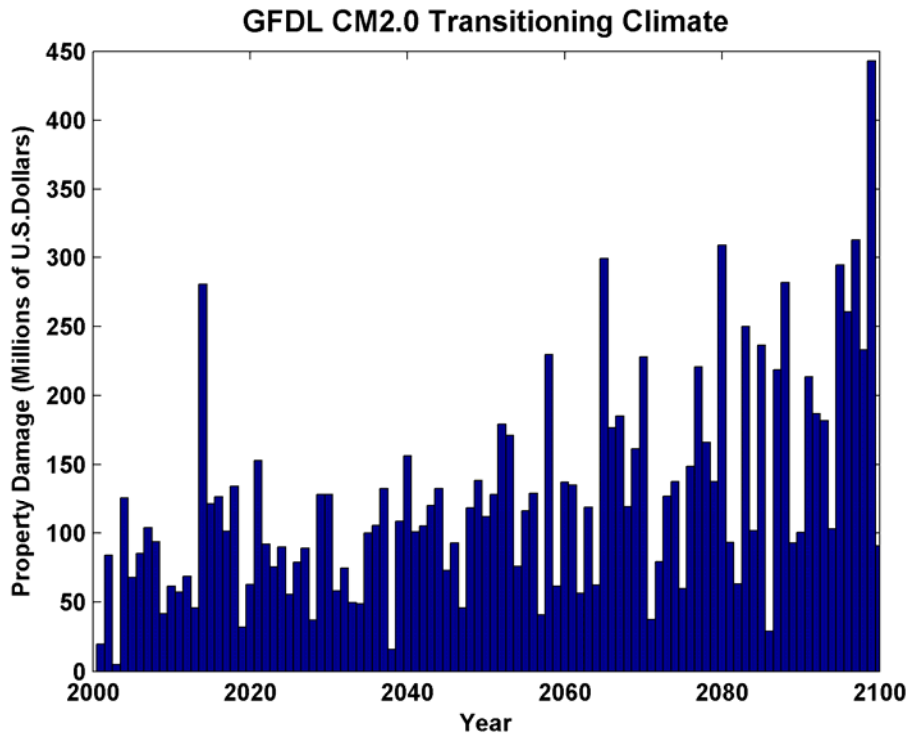
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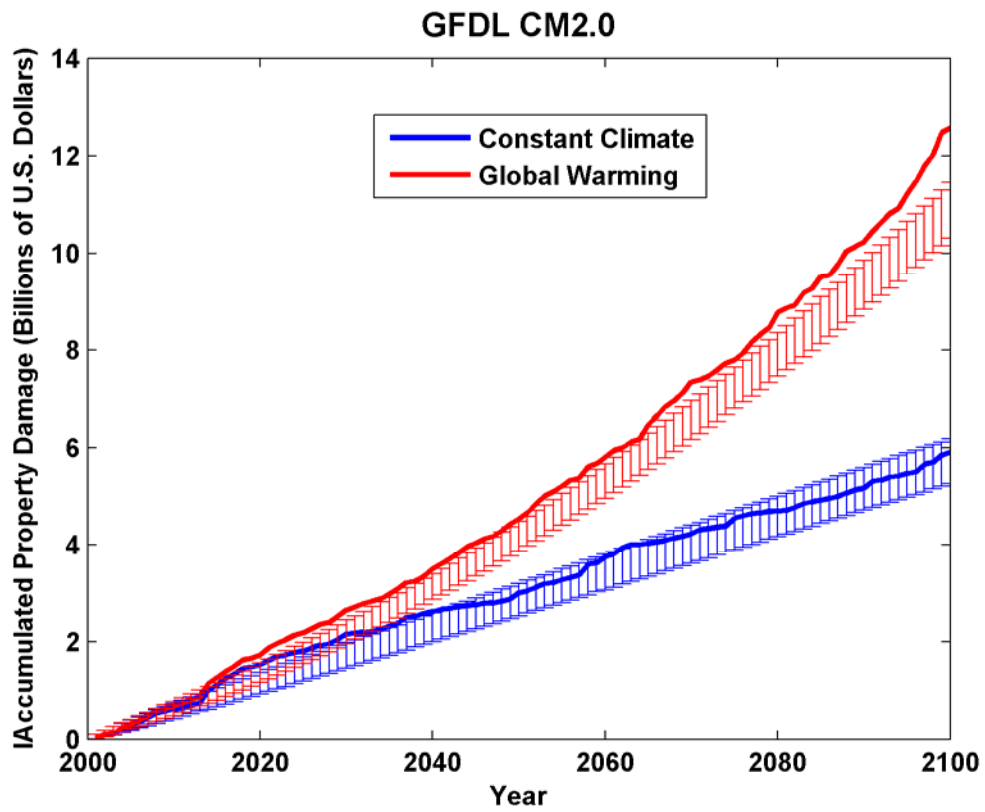
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410 Figure 4: Same as Figure 3 but for a climate transitioning linearly from its state at the end of the 20<sup>th</sup> century to its state at  
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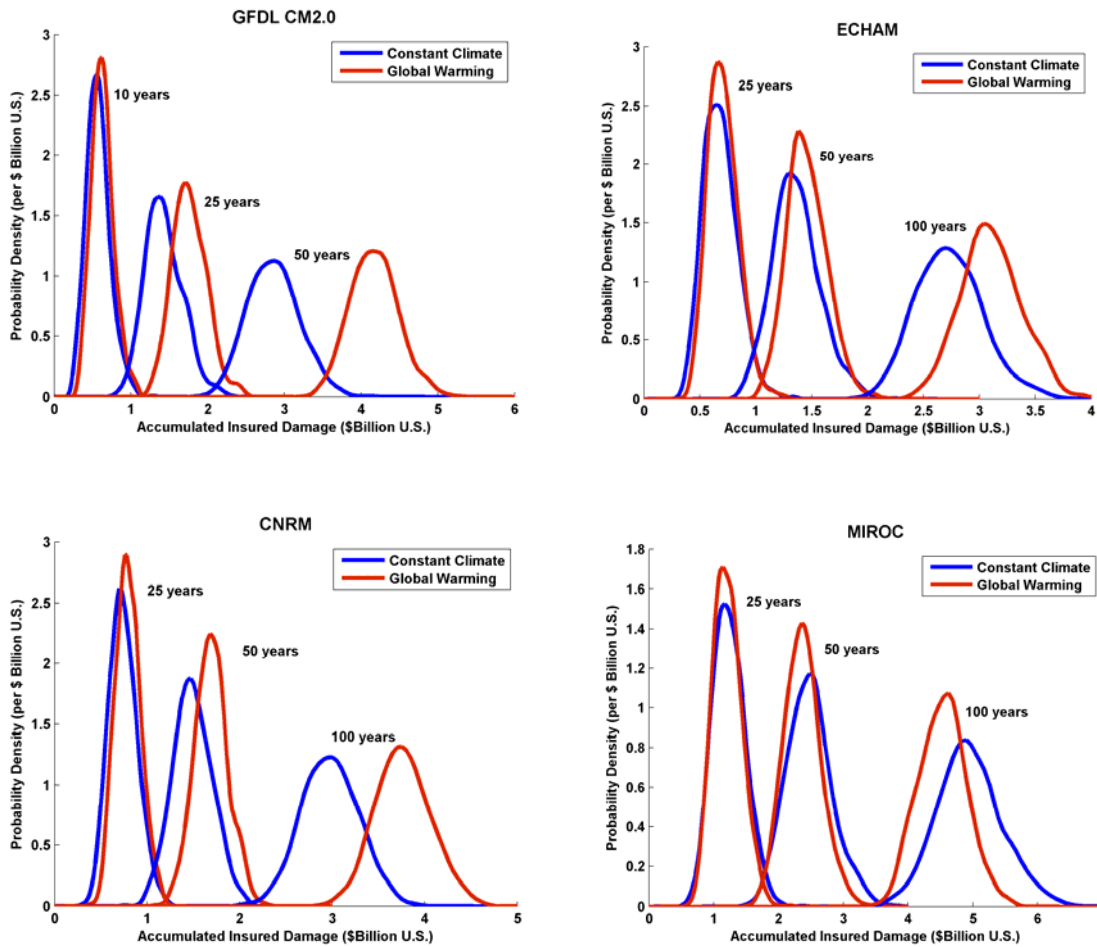


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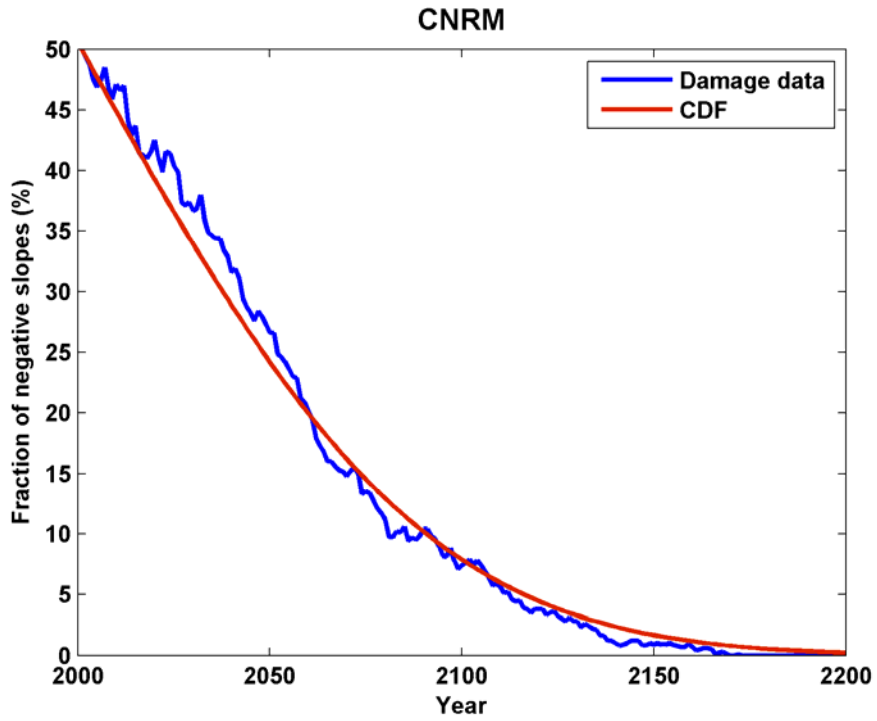
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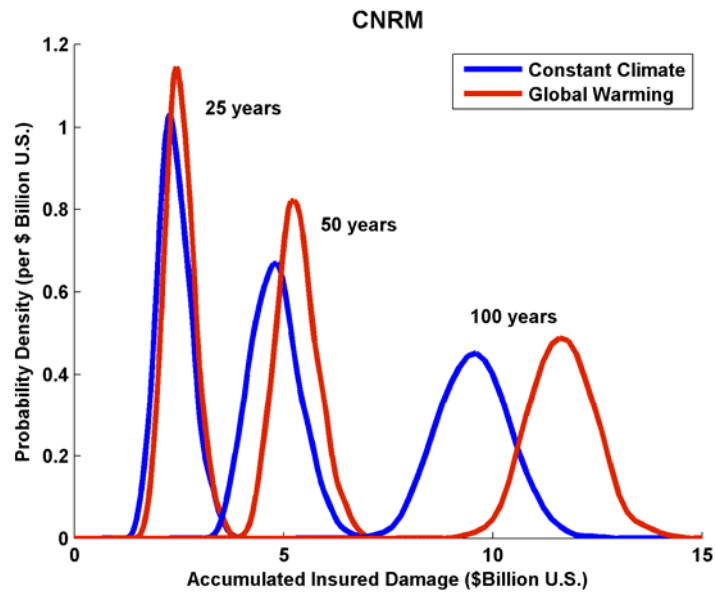
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